

Lecture 8

Longhai Li, October 5, 2021

Def of $E(X)$:

1. X is discrete.

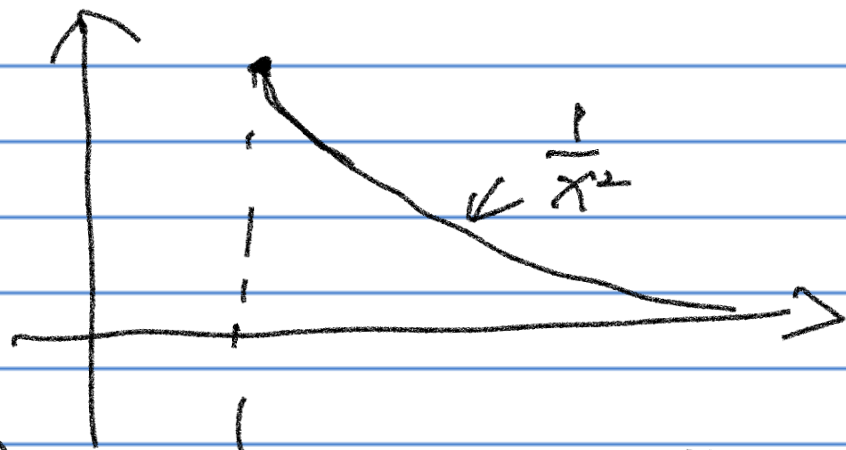
$$E(g(X)) = \sum_{\text{all possible } X} g(x) \cdot p(x)$$

2. X is continuous

$$E(g(X)) = \int_{-\infty}^{+\infty} g(x) f(x) dx$$

Example of non-finite Expectation:

$$f(x) = \frac{1}{x^2}, \text{ for } x \geq 1$$



$$\int_1^{+\infty} f(x) dx = \int_1^{+\infty} \frac{1}{x^2} dx = -\frac{1}{x} \Big|_1^{+\infty} \\ = 0 + 1 = 1$$

Integrals of power functions



$$\int_c^{+\infty} \frac{1}{x} = +\infty$$

$$\int_c^{+\infty} \frac{1}{x^{1+\delta}} \leq +\infty$$

for $\delta > 0$

$$\int_c^{+\infty} \frac{1}{x^{1+\delta}} = +\infty$$

for $\delta \leq 0$

$$f(x) = \frac{1}{x^2}, \text{ for } x \geq 1$$

$$E(x) = \int_1^{\infty} x \cdot f(x) dx$$

$$= \int_1^{\infty} x \cdot \frac{1}{x^2} dx$$

$$= \int_1^{\infty} \frac{1}{x} dx = \infty.$$

$$= \log x \Big|_1^{\infty} = \infty - 0 = \infty$$

Linearity of E .

$$E(k_1 g_1(x) + k_2 g_2(x))$$

$$= k_1 E(g_1(x)) + k_2 E(g_2(x)).$$

Pf:

Suppose X is continuous.

~~QED~~

$$L(f+S) = \int_{-\infty}^{\infty} (k_1 f_1(x) + k_2 f_2(x)) dx$$

$$= k_1 \int_{-\infty}^{\infty} f_1(x) dx + k_2 \int_{-\infty}^{\infty} f_2(x) dx$$

$$= k_1 E(f_1(x)) + k_2 E(f_2(x))$$

Thm: $X \geq 0$ is a r.v. $F(x)$ is the
C.D.F. of X . Then

$$E(X) = \int_0^{+\infty} [1 - F(x)] dx$$

↑
survival function

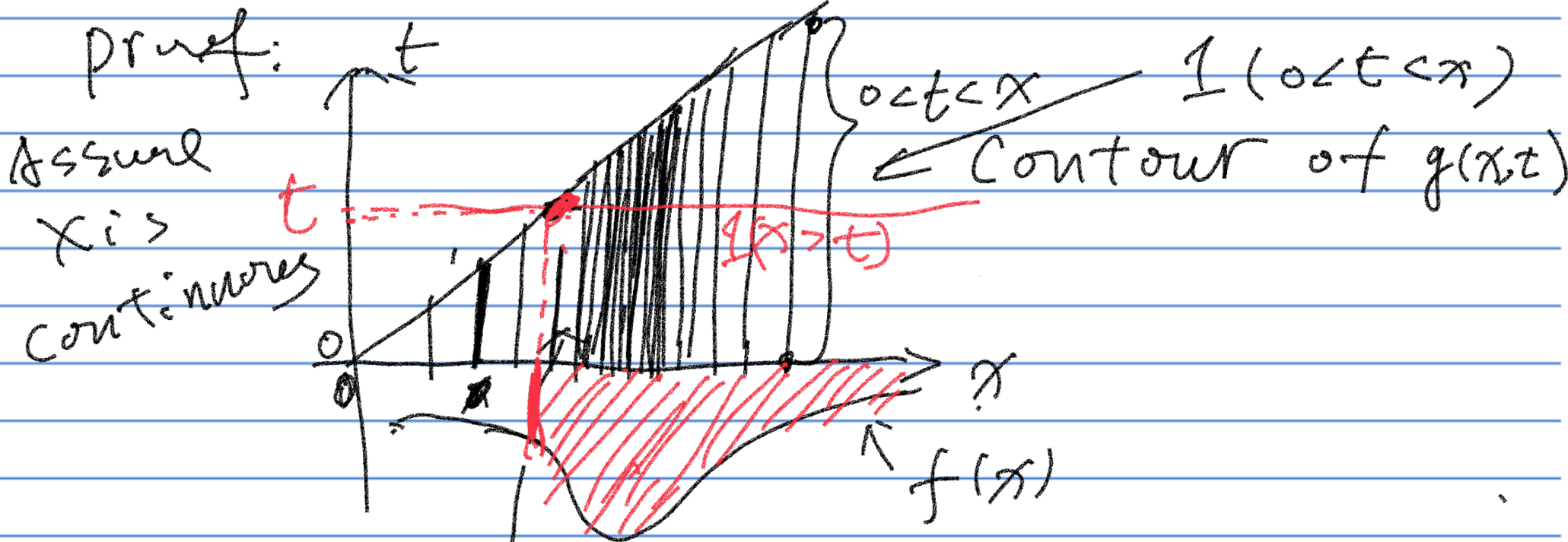
General. $X^+ = \max(X, 0)$

$X^- = \max(-X, 0)$

$$X = X^+ - X^-$$

$$E(X) = E(X^+) - E(X^-)$$





$$g(x,t) = f(x) \cdot \boxed{I(0 < t < x)}$$

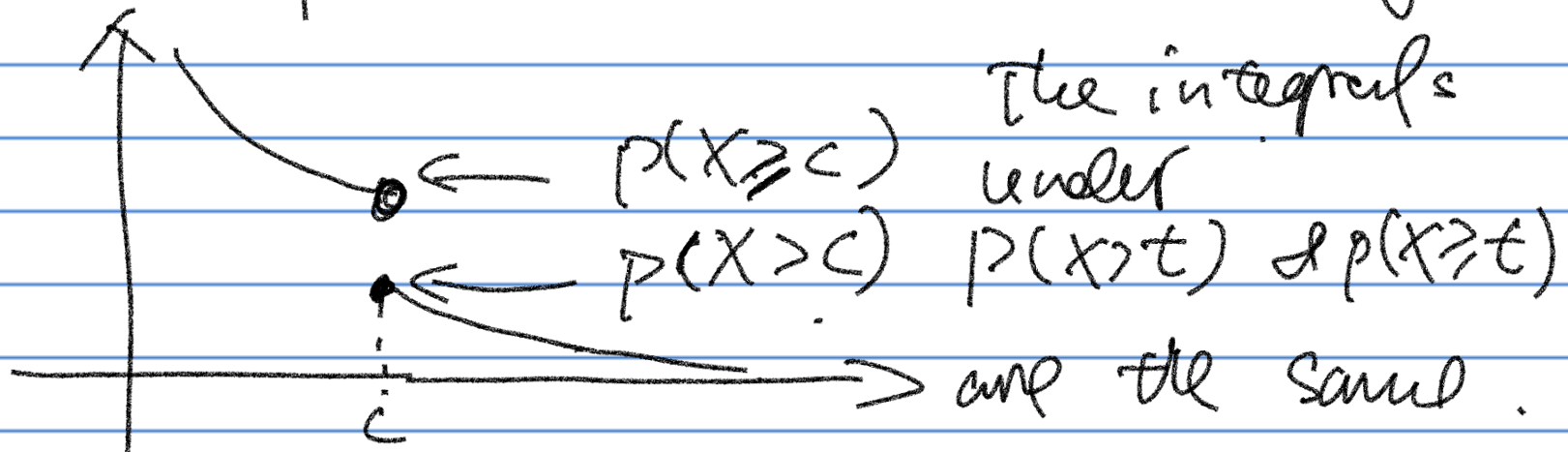
$$\begin{aligned}
 \int_{-\infty}^{+\infty} x f(x) dx &= \int_{-\infty}^{+\infty} f(x) \left(\int_0^x 1 dt \right) dx \\
 \underbrace{\hspace{10em}}_{\text{Fubini}} &= \int_{-\infty}^{+\infty} f(x) \int_{-\infty}^{+\infty} \mathbb{1}(0 < t < x) dt dx \\
 &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x) \mathbb{1}(0 < t < x) dt dx \\
 &\stackrel{\text{Fubini Theorem.}}{=} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x) \mathbb{1}(0 < t < x) dx dt \\
 &= \int_{-\infty}^{+\infty} \int_t^{+\infty} f(x) dx dt \\
 &= \int_{-\infty}^{+\infty} [1 - F(t)] dt
 \end{aligned}$$

A note:

$$E(X) = \int_0^{+\infty} P(X > t) dt$$

$$= \int_0^{+\infty} P(X \geq t) dt$$

$P(X > t) = P(X \geq t)$ for almost everywhere



Sec 1.9. Special Exp.

Mean: $E(X) = \mu$

Variance:

$$V(X) = E((X - \mu)^2)$$

$$= E(X^2 - 2\mu X + \mu^2)$$

$$= E(X^2) - \mu^2$$

$$= E(X^2) - [E(X)]^2$$

1) $E(X), V(X)$ may be non-finite.

2) $V(X) > 0 \Rightarrow E(X^2) \geq [E(X)]^2$

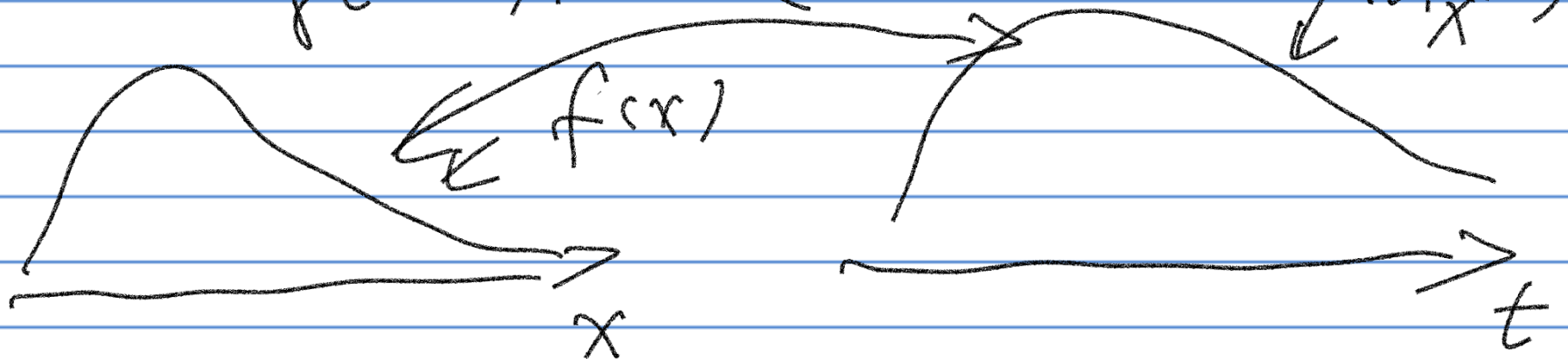
Def of M. G. F.:

$$M_x(t) = E(e^{tx})$$

$f_t(x)$

t is a fixed value.

Let $f_t(x) = e^{tx}$



Remarks:

1). $M_X(t)$ may not exist.

$M_X(t) = E(e^{tx})$ may be $= +\infty$.

$M_X(t)$ exists for $t \in (-h, h)$

for some $h \iff E(X^k) < +\infty$

for all $k = 0, 1, \dots$

Moments

$$2) M_X^{(k)}(0) = E(X^k) \quad \text{moments.}$$

$$\text{Pf: } k=1.$$

$$M_X(t) = E(e^{tX})$$

$$M_X^{(1)}(t) = E(e^{tX} \cdot X)$$

$$M_X^{(1)}(0) = E(X)$$

$$\begin{aligned} M_X^{(2)}(t) &= E(e^{tX} \cdot X \cdot X) \\ &= E(e^{tX} \cdot X^2) \end{aligned}$$

$$M^{(2)}(0) = E(X^2)$$

$$M^{(k)}(t) = E(e^{tX} \cdot X^k)$$

$$M^{(k)}(0) = E(X^k)$$

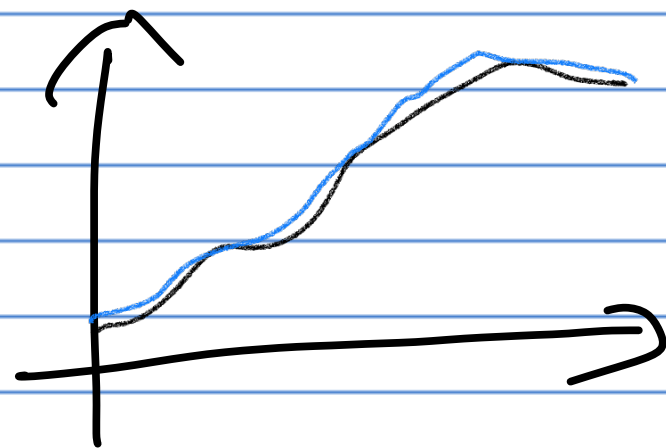
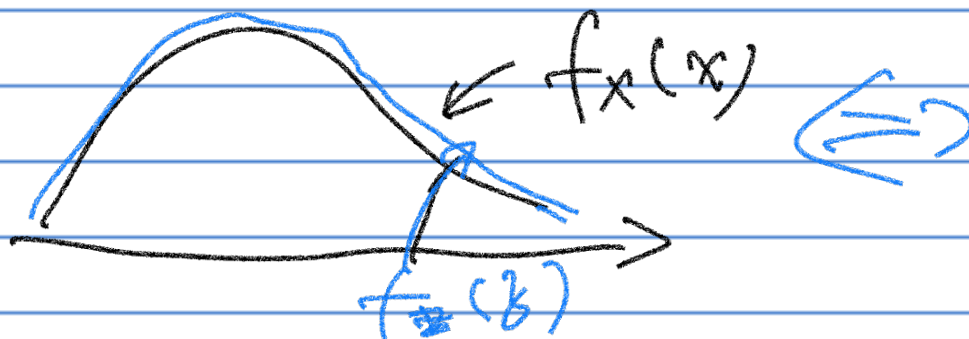
Thm: (Thm 1.9.1.)

Let X & Y be two r.v.

$$F_X(x) = F_Y(y) \quad \text{for all } x \text{ & } y$$

$$\Leftrightarrow M_X(t) = M_Y(t) \quad \text{for } t \in (-h, h)$$

for some $h > 0$.



Characteristic Function (always exist)

i is the imaginary number. ($i^2 = -1$)

$$C_X(t) = E(e^{itX})$$

$$= E(\cos(tX) + i \sin(tX))$$

$$= E(\cos(tX)) + i E(\sin(tX))$$

$$< +\infty, |\cos(tX)| < 1, |\sin(tX)| < 1$$

Section 1.10

Important Inequalities

Markov Inequality

Then: $X \geq 0$, $P(X \geq c) \leq \frac{E(X)}{c}$
for any $c > 0$.

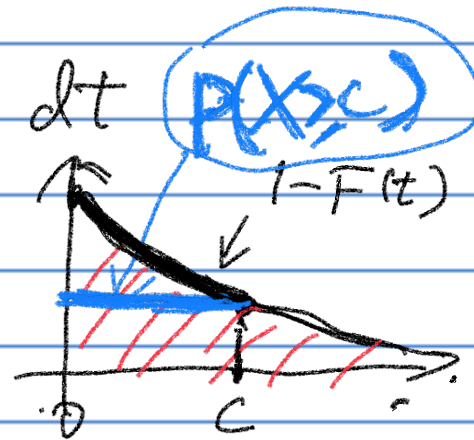
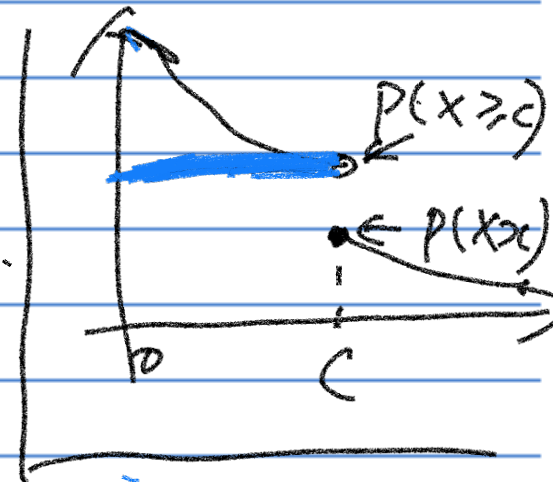
Pf: $E(X) = \int_0^{\infty} P(X \geq t) dt$

$$\geq \int_0^c P(X \geq t) dt$$

$$\geq P(X \geq c) \cdot c$$

$P(X \geq t) \geq P(X \geq c)$, for $t \leq c$

$$P(X \geq c) \leq \frac{E(X)}{c}$$



Chebyshev's Inequality:

Suppose $V(X)$ exists, $\mu = E(X)$

$$P(|X - \mu| \geq c) \leq \frac{V(X)}{c^2}$$

Pf: LHS

$$= P(|X - \mu|^2 \geq c^2)$$

$$\leq \frac{E(|X - \mu|^2)}{c^2}$$

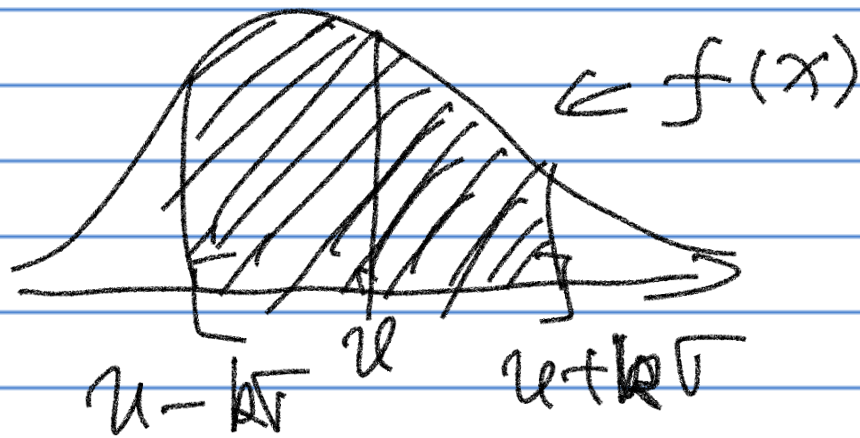
$$\leq \frac{V(X)}{c^2}$$

, by Markov's Ineq.

Another form of Chebyshev's Ineq.

$$c = k \cdot \sigma = k \cdot \sqrt{V(x)}$$

$$P(|X - \mu| > k \cdot \sigma) \leq \frac{V(x)}{k^2 \cdot \sigma^2} = \frac{1}{k^2}$$



$$\sigma = \text{sd}(x)$$

$$k=3, \frac{1}{k^2} = \frac{1}{9}$$

$$k=4, \frac{1}{k^2} = \frac{1}{16}$$

Sometimes useful
in practice.

Thm:

If $E(X^m) < +\infty$ exists

then $E(X^k) < +\infty$ for all $k \leq m$.

pf:

$$E(|X|^k) = E(\overset{\leq 1}{|X|^k} \mathbb{1}(|X| \leq 1)) \leq E(\mathbb{1}(|X| \leq 1)) + E(|X|^k \mathbb{1}(|X| > 1)) \leq E(|X|^m \mathbb{1}(|X| > 1))$$

$$\leq P(|X| \leq 1) + E(|X|^m) < \infty$$

$E(X^k)$ is finite

exist. finite. $< +\infty$, §. 1. 2. 1.

non-existent, non-finite.

$\left\{ \begin{array}{l} +\infty \end{array} \right.$

undefined: $+\infty - +\infty$