

Lecture 14

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plan:

Sec 2-6, 2-7

1. extension to ≥ 3 random variables.
2. Linear combination (Sec 2-8)

≥ 3 Random variables (S2-6)

Continuous Random Variables

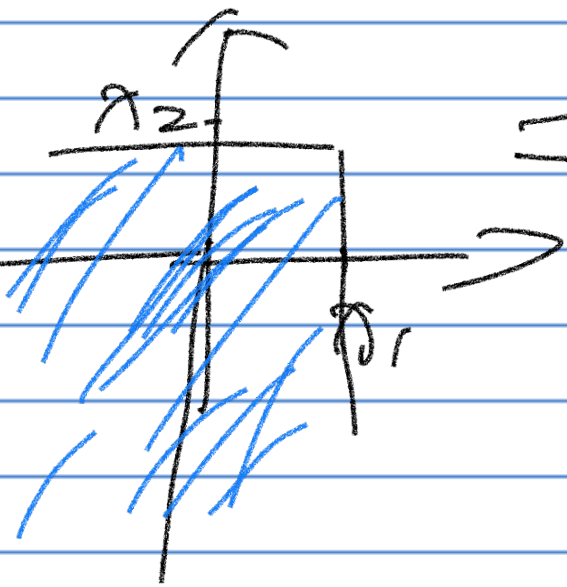
$f(x_1, x_2, \dots, x_n)$ is a joint P.D.F. of X_1, \dots, X_n if

$$\int \dots \int_A f(x_1, \dots, x_n) \underline{dx_1 \dots dx_n}$$

$$= P((X_1, \dots, X_n) \in A)$$

Joint C.D.F.

$$F(x_1, \dots, x_n) = P(X_1 \leq x_1, \dots, X_n \leq x_n)$$



$$= \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} \dots \int_{-\infty}^{x_n} f(x_1, \dots, x_n) \cdot dx_1 \dots dx_n$$

Marginal P.D.F.
 $s < n$

$$f(x_1, \dots, x_s)$$

$$= \int_{x_{s+1}} \dots \int_{x_n} f(x_1, \dots, x_s, x_{s+1}, \dots, x_n) dx_{s+1} \dots dx_n$$

Conditional P.D.F.

$$f(x_{s+1}, \dots, x_n | x_1, \dots, x_s) = \frac{f(x_1, \dots, x_n)}{f(x_1, \dots, x_s)}$$

Generally, think X_1, X_2 as random vectors

Joint M.G.F.

$$M(t_1, \dots, t_n) = E\left(e^{\underline{t_1 X_1 + \dots + t_n X_n}}\right)$$

Expectation

$$E(g(X_1, \dots, X_n)) = \int \int g(x_1, \dots, x_n) \cdot f(x_1, \dots, x_n) \cdot dx_1 \dots dx_n$$

Independence of X_1, \dots, X_n :

We say X_1, X_2, \dots, X_n are indep

if

$$f(x_1, x_2, \dots, x_n) \\ = f_{X_1}(x_1) \dots f_{X_n}(x_n)$$

for all x_1, x_2, \dots, x_n

Def of Random Sample (I.I.D.)

X_1, X_2, \dots, X_n are independent

and identically distributed, i.e.

$$f(x_1, x_2, \dots, x_n) = \prod_{i=1}^n f_i(x_i)$$

Here $f(x_i)$ is a P.D.F. of X_i .

Thm: M.G.F. of Sum of Random Sample

Suppose X_1, \dots, X_n are IID,

let $T = \sum_{i=1}^n X_i$ (sample total)

then $M_T(t) = [M_{X_i}(t)]^n$ ✓

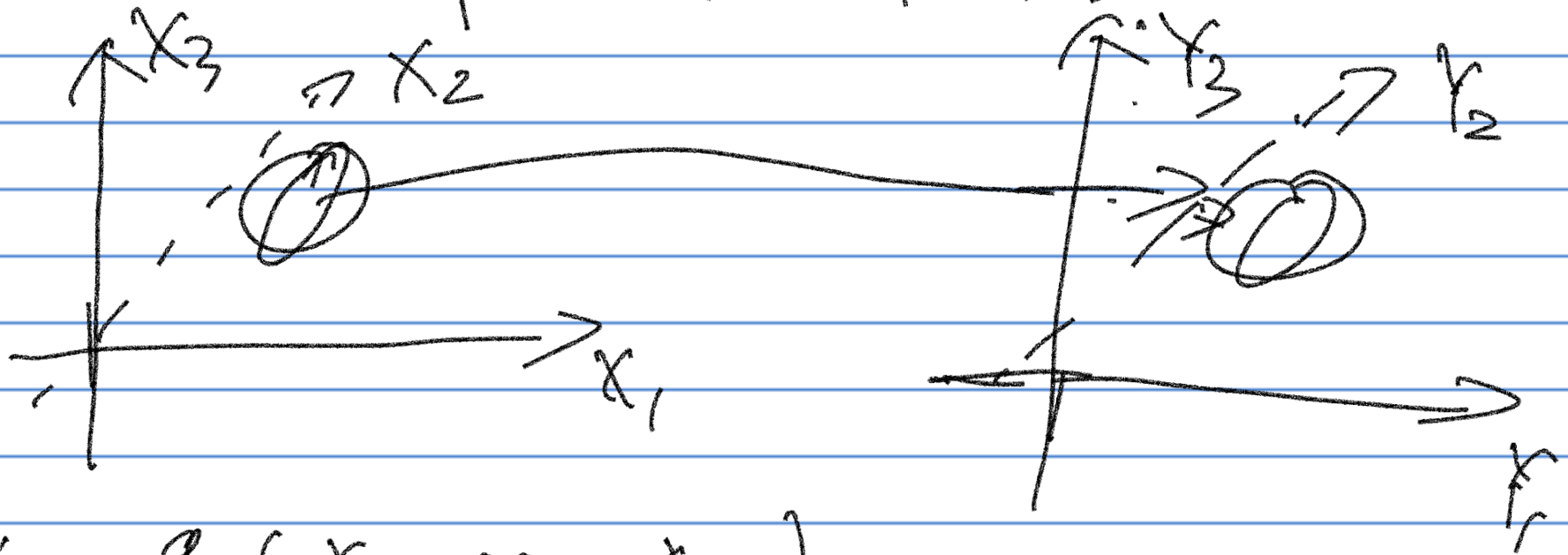
$$\text{Pf } M_T(s) = E(e^{sT}) = E(e^{s \sum_{i=1}^n X_i})$$

$$= E(e^{sX_1} \cdot e^{sX_2} \cdot \dots \cdot e^{sX_n})$$

$$= E(e^{sX_1}) \cdot \dots \cdot E(e^{sX_n})$$

$$= M_{X_1}(s) \cdot \dots \cdot M_{X_n}(s) = [M_{X_i}(s)]^n$$

Sec 2.7. Transformation of ≥ 3 r.v.



$$Y_1 = g_1(x_1, \dots, x_n)$$

$$\left\{ \begin{array}{l} Y_1 = g_1(x_1, \dots, x_n) \\ Y_n = g_n(x_1, \dots, x_n) \end{array} \right.$$

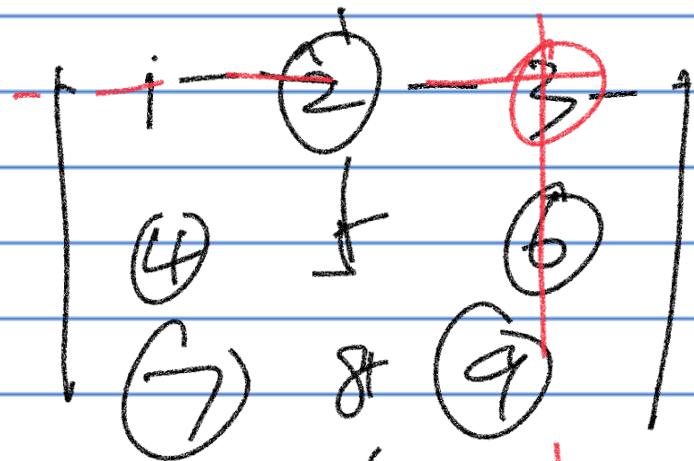
1-1

$$\begin{cases} x_1 = t_1(y_1, \dots, y_n) \\ \vdots \\ x_n = t_n(y_1, \dots, y_n) \end{cases}$$

$$f_y(y_1, \dots, y_n) = f_x(x_1, \dots, x_n) \cdot \bar{J}$$

$$\bar{J} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \dots & \frac{\partial x_1}{\partial y_n} \\ \frac{\partial x_2}{\partial y_1} & \dots & \frac{\partial x_2}{\partial y_n} \\ \vdots & \dots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \dots & \frac{\partial x_n}{\partial y_n} \end{vmatrix} \dagger$$

\bar{J} Jacobian matrix.



$$= 1 \cdot \begin{vmatrix} 5 & 6 \\ 8 & 9 \end{vmatrix} - 2 \cdot \begin{vmatrix} 4 & 6 \\ 7 & 9 \end{vmatrix}$$

$$+ 3 \cdot \begin{vmatrix} 4 & 5 \\ 7 & 8 \end{vmatrix}$$

$$\begin{vmatrix} \textcircled{x_1} & x & x \\ 0 & \textcircled{x_2} & x \\ 0 & 0 & \textcircled{x_3} \end{vmatrix}$$

$$= x_1 x_2 x_3$$

Linear Combination of Random Variables

(X_1, \dots, X_n) is a random vector.

$$\text{let } T = \sum_{i=1}^n a_i X_i = a_1 X_1 + a_2 X_2 + \dots + a_n X_n$$

Thm: Suppose $E(|X_i|) < \infty$ (No assumption of independence)

$$T = \sum_{i=1}^n a_i X_i,$$

Then $E(T) = \sum_{i=1}^n a_i E(X_i)$. (linearity of E .)

pf: $E(T) = \int \dots \int \left(\sum_{i=1}^n a_i x_i \right) f(x_1, \dots, x_n) dx_1 \dots dx_n$

$$= \sum_{i=1}^n \int \dots \int \underbrace{a_i x_i}_{\text{circled}} \underbrace{f(x_1, \dots, x_n)}_{\text{underlined}} dx_1 \dots dx_n$$

$$= \sum_{i=1}^n a_i \cdot E(X_i)$$

Thm:

$$T = \sum_{i=1}^n a_i X_i, \quad W = \sum_{j=1}^m b_j Y_j$$

$$\text{Cov}(T, W) = \sum_{i=1}^n \sum_{j=1}^m a_i b_j \text{Cov}(X_i, Y_j)$$

Prf: Let $\mu_i = E(X_i)$, $\sigma_j = E(Y_j)$

$$\text{Cov}(T, W) = E\left((T - E(T)) \cdot (W - E(W)) \right)$$

$$= E\left(\left[\sum_{i=1}^n a_i (X_i - \mu_i) \right] \cdot \left[\sum_{j=1}^m b_j (Y_j - \sigma_j) \right] \right)$$

$$= E\left(\sum_{i=1}^n \sum_{j=1}^m a_i b_j (X_i - \mu_i) (Y_j - \sigma_j) \right)$$

$$= \sum_{i=1}^n \sum_{j=1}^m a_i b_j \text{Cov}(X_i, Y_j)$$

Thm:

$$T = \sum_{i=1}^n a_i X_i$$

$$V(T) = \sum_{i=1}^n a_i^2 V(X_i) + 2 \sum_{i < j} a_i a_j \text{Cov}(X_i, X_j)$$

Pr:

$$V(T) = \text{Cov}(T, T)$$

$$= \sum_{i=1}^n \sum_{j=1}^n a_i a_j \text{Cov}(X_i, X_j)$$

$$= \sum_{i=1}^n a_i^2 V(X_i) + \left(2 \sum_{i < j} a_i a_j \text{Cov}(X_i, X_j) \right)$$

Thm: ^{assume} X_1, \dots, X_n are independent.

$$\text{Let } T = \sum_{i=1}^n a_i X_i$$

$$\text{Then } V(T) = \sum_{i=1}^n a_i^2 V(X_i)$$

"Linearity" of Variance when
 X_i 's are indep.

Pf: When $X_i \perp X_j$, $\text{Cov}(X_i, X_j) = 0$.

Example:

$$V(\bar{X}_1 - \bar{X}_2) = V(\bar{X}_1) + V(\bar{X}_2)$$

if \bar{X}_1 and \bar{X}_2 indep

$$V(\bar{X}_1) = V\left(\frac{\sum_{i=1}^n X_i}{n}\right)$$

$$= \frac{1}{n^2} \cdot V\left(\sum_{i=1}^n X_i\right)$$

$$= \frac{1}{n^2} \cdot \sum_{i=1}^n V(X_i) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$$

where $V(X_i) = \sigma^2$

Example:

X_1, \dots, X_n is a random sample, i.e.,

X_1, \dots, X_n ^{are} **IID**.

let $\mu = E(X_i)$, $\sigma^2 = V(X_i)$.

let $\bar{X} = \frac{X_1 + \dots + X_n}{n}$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

Then: $E(\bar{X}) = \mu$, $E(S^2) = \sigma^2$, $V(\bar{X}) = \frac{\sigma^2}{n}$

Pf:

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i)$$

$$= \frac{1}{n} \cdot n \cdot \mu = \mu, \quad V(\bar{X}) = \frac{\sigma^2}{n}$$

$$E(S^2) = E\left(\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2\right)$$

$$= E\left(\frac{1}{n-1} \sum_{i=1}^n (X_i^2 - 2\bar{X} \cdot X_i + \bar{X}^2)\right)$$

$$= E\left(\frac{1}{n-1} \left[\sum_{i=1}^n X_i^2 - n\bar{X}^2\right]\right)$$

$$= \frac{1}{n-1} \left[n(\mu^2 + \sigma^2) - n \cdot \left(\mu^2 + \frac{\sigma^2}{n}\right) \right]$$

$$= \frac{1}{n-1} E[(n-1)\sigma^2] = \sigma^2$$