

Lecture 17

Longhai Li, Nov 4, 2021

Plan:

1) Map on Gauss

2) Normal distribution.

Special Gamma.

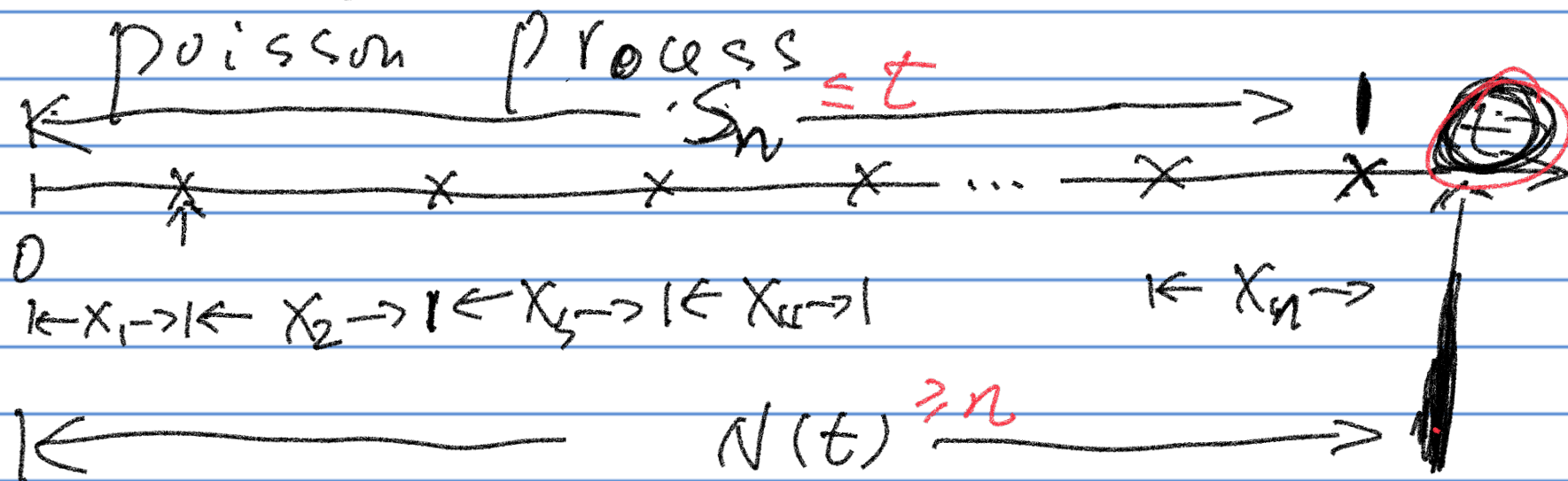
$$1) \alpha = 1, \beta = 1$$

$$f(x) = \frac{e^{-x} x^{\alpha-1}}{\Gamma(\alpha)} = e^{-x} \text{ for } x > 0$$

$$\text{Gamma}(\alpha=1, \beta=1) = \exp(-x)$$

$$\text{Gamma}(\alpha=1, \beta) = \exp\left(-\frac{x}{\beta}\right)$$

2) Erlang distribution $\alpha = n$.



$N(t) = \#$ of events btw $[0, t]$

$$S_n = X_1 + X_2 + \dots + X_n$$

$N(t)$ takes values $0, 1, 2, \dots$

S_n is continuous, $S_n \geq 0$

This process is called Poisson process
with rate λ if

$$N(t) \sim \text{poisson}(\lambda t) \quad \checkmark$$

and something more,

$$E(N(t)) = \lambda t$$

$$E(N(1)) = \lambda$$

Thm:

$$N(t) \sim \text{Poisson}(\lambda t)$$

$$\Leftrightarrow S_n = X_1 + \dots + X_n \sim \text{Gamma}(n, \beta = \frac{1}{\lambda}).$$

X_1, \dots, X_n are ~~IID~~ $\exp(\lambda)$

Pf:

$$F_{S_n}(t) = P(S_n \leq t) = P(N(t) \geq n)$$

$$P(S_n \leq t) = P(N(t) \geq n)$$

$$= \sum_{k=n}^{+\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!}$$

$$F'_{S_n}(t) = \left[\frac{d}{dt} \sum_{k=n}^{+\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right]$$

$$= \sum_{k=n}^{+\infty} \frac{e^{-\lambda t} (\lambda t)^k + e^{-\lambda t} (\lambda t)^{k-1} \cdot (-1)}{k!}$$

$(k-1)!$

$$= \lambda \cdot \left[\sum_{k=n}^{\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!} + \sum_{k=n}^{\infty} \frac{e^{-\lambda t} (\lambda t)^{k-1}}{(k-1)!} \right]$$

$$= \lambda \left[\sum_{k=n}^{\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!} + \sum_{k=n-1}^{\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right]$$

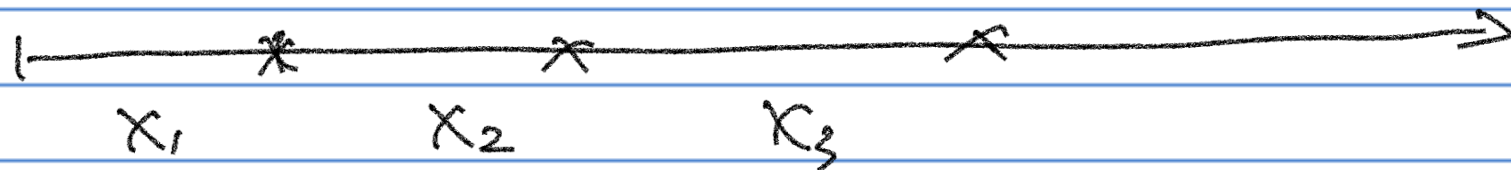
$$= \lambda \frac{e^{-\lambda t} (\lambda t)^{n-1}}{(n-1)!}$$

This is the P.D.F. of Gamma($n, \beta = \frac{1}{\lambda}$)

$$S_n \sim \text{Gamma}(n, \beta = \frac{1}{\lambda})$$

Erlang distribution

$$S_1 = X_1 \sim \text{Gamma}(1, \frac{1}{\lambda}) = \exp(\lambda)$$



$$X_1, X_2, \dots, X_n \stackrel{\text{IID}}{\sim} \exp(\lambda)$$

(1)
 $\text{Gamma}(1, \frac{1}{\lambda})$

$$S_n = X_1 + \dots + X_n$$

Sum of n IID $\exp(\lambda)$ r.v.

Thm: additivity of Gamma

If $X_i \sim \text{Gamma}(\underline{\alpha}_i, \beta)$,

X_1, \dots, X_n are independent.

then $\underline{X_1 + \dots + X_n} \sim \text{Gamma}(\sum_{i=1}^n \alpha_i, \beta)$

Prf: $M_{X_i}(t) = \left(1 - \frac{t}{\beta}\right)^{-\alpha_i}$

$$M_{X_1 + \dots + X_n}(t) = \prod_{i=1}^n M_{X_i}(t) = \prod_{i=1}^n \left(1 - \frac{t}{\beta}\right)^{\alpha_i}$$

$$= \left(1 - \frac{t}{\beta}\right)^{\sum_{i=1}^n \alpha_i}$$

This is the M.G.F. of $\text{Gamma}(\sum_{i=1}^n \alpha_i, \beta)$

Additivity of Binomial.

If $X_i \sim \text{Bin}(n_i, p)$, indep.

then $S_n = X_1 + \dots + X_n \sim \text{Bin}(\sum_{i=1}^k n_i, p)$.

pf: $M_{X_i}(t) = (q + p e^t)^{n_i}$

$$\begin{aligned} M_{S_n}(t) &= \prod_{i=1}^k M_{X_i}(t) \\ &= (q + p e^t)^{\sum_{i=1}^k n_i} \end{aligned}$$

Addition of poisson(λ).

If $X_i \sim \text{poisson}(\lambda_i)$, indep

for $i=1, \dots, n$

then $S_n = X_1 + \dots + X_n \sim \text{poisson}(\sum_{i=1}^n \lambda_i)$

pf: $M_{X_i}(t) = e^{\lambda_i(e^t - 1)}$

$$M_{S_n}(t) = e^{\sum_{i=1}^n \lambda_i (e^t - 1)}$$

Beta distribution.

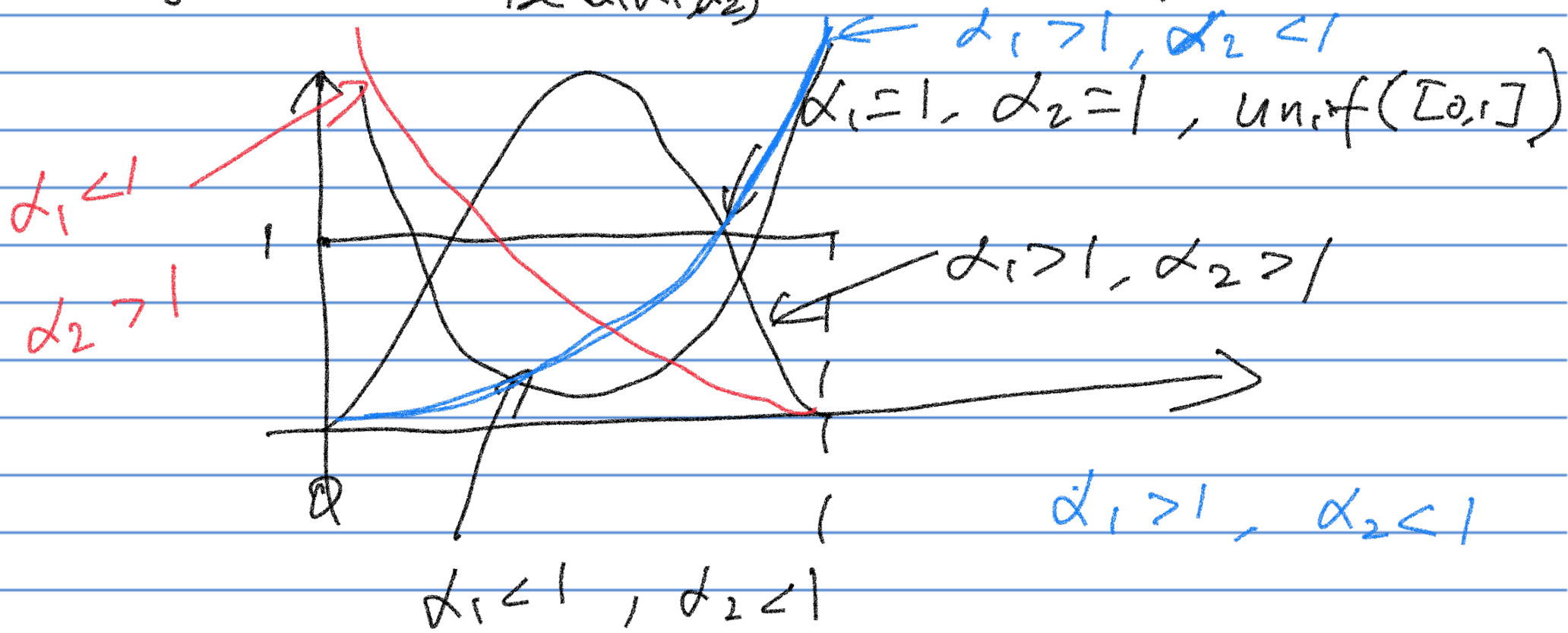
$$X \sim \text{Gamma}(\alpha_1, 1)$$

$$Y \sim \text{Gamma}(\alpha_2, 1)$$

$$W = \frac{X}{X+Y} \sim \text{Beta}(\alpha_1, \alpha_2)$$

$$W \in [0, 1]$$

$$f(\omega) = \frac{\omega^{\alpha_1 - 1} (1 - \omega)^{\alpha_2 - 1}}{\text{Beta}(\alpha_1, \alpha_2)}, \quad \text{for } \omega \in [0, 1]$$



$$E(X) = \frac{\alpha_1}{\alpha_1 + \alpha_2}$$

$$V(X) = \frac{\alpha_1 \cdot \alpha_2}{(\alpha_1 + \alpha_2 + 1)(\alpha_1 + \alpha_2)^2}$$

A remark: $X \sim \text{Gamma}(\alpha_1, \beta)$, $\frac{X}{\beta} \sim \text{Gamma}(\alpha_1, 1)$
 $Y \sim \text{Gamma}(\alpha_2, \beta)$, $\frac{Y}{\beta} \sim \text{Gamma}(\alpha_2, 1)$

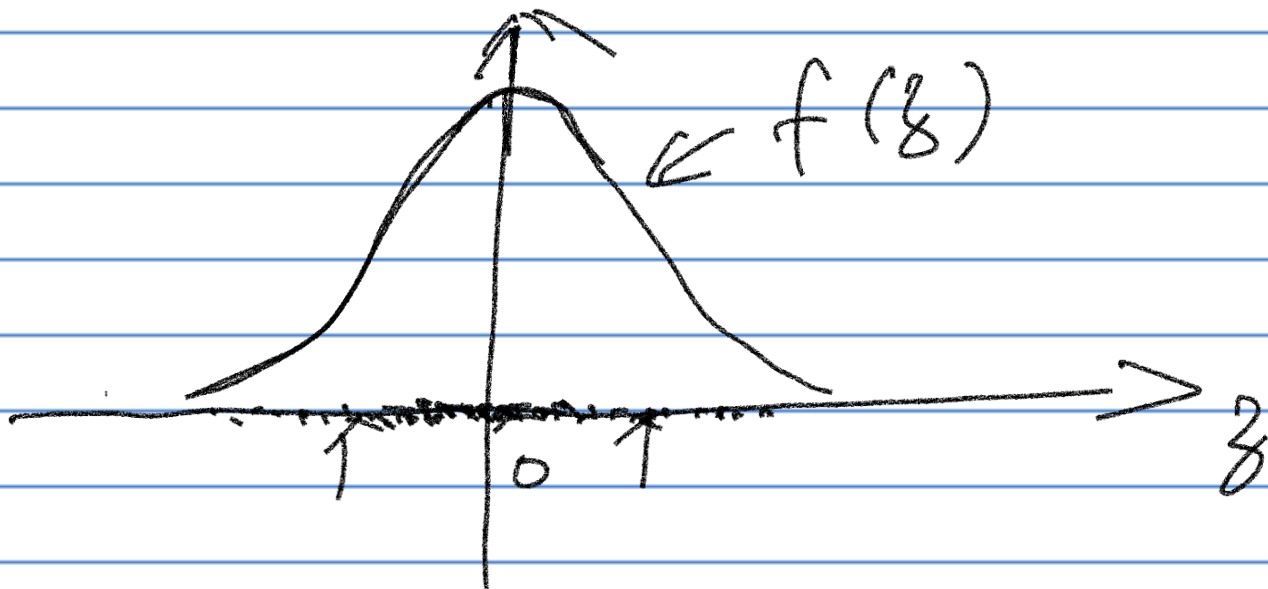
$$W = \frac{X}{X+Y} = \frac{\frac{X}{\beta}}{\frac{X}{\beta} + \frac{Y}{\beta}} \sim \text{Beta}(\alpha_1, \alpha_2)$$

Normal distribution

$Z \sim N(0, 1)$:

P.D.F.

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < +\infty$$



$$E(z) = 0$$

$$V(z) = E(z^2)$$

$$= \int_{-\infty}^{+\infty} z^2 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$$

$$= 2 \frac{1}{\sqrt{2\pi}} \int_0^{+\infty} z^2 e^{-\frac{z^2}{2}} dz \quad z = \sqrt{x}$$

$$= 2 \cdot \frac{1}{\sqrt{2\pi}} \int_0^{+\infty} x e^{-\frac{x}{2}} d\sqrt{x}$$

...

$$= 1$$

M.G.F.

$$M_Z(t) = E(e^{tz})$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{tz} e^{-\frac{z^2}{2}} dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{1}{2} [z^2 - 2tz + t^2 - t^2]} dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{1}{2} [z-t]^2} e^{+\frac{1}{2}t^2} dz$$

$$= e^{\frac{1}{2}t^2} \cdot \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{1}{2}(z-t)^2} dz, \quad z=t \rightarrow x$$

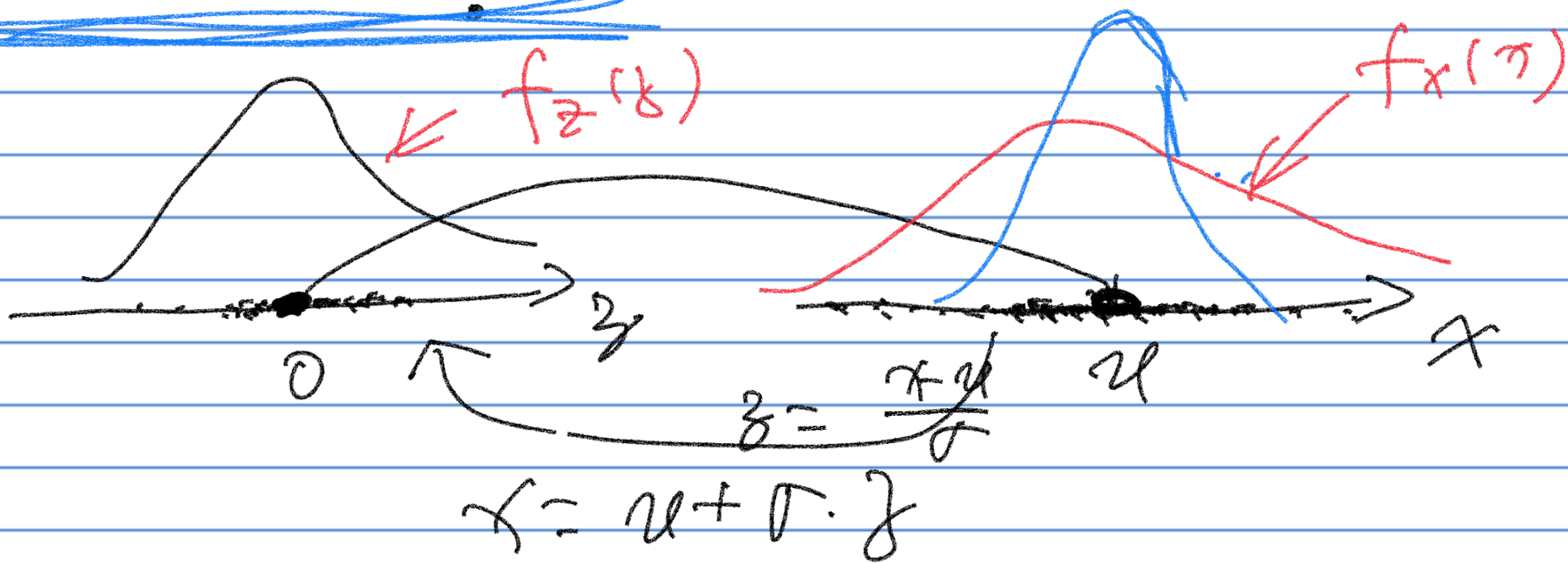
$$= e^{\frac{1}{2}t^2} \left(\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{-\frac{1}{2}x^2} dx \right) \rightarrow 1$$

$$= e^{\frac{1}{2}t^2}$$

$$F_z(z) = \int_{-\infty}^z f(z) dz, \text{ no closed form}$$

z	$F(z)$
-3.0	0.00 . . .
-5.1	

$$X = \mu + \sigma Z, \quad Z \sim N(0, 1)$$



$$X \sim N(\mu, \sigma^2)$$

D.D.F. of X

$$Z = \frac{x - \mu}{\sigma}$$

$$f_X(x) = f_Z\left(\frac{x - \mu}{\sigma}\right) \left| \frac{dz}{dx} \right| \cdot \sigma$$

$$= f_Z\left(\frac{x - \mu}{\sigma}\right) \left(\frac{1}{\sigma} \right)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{(x - \mu)^2}{2\sigma^2}} \cdot \frac{1}{\sigma}$$

$$\begin{aligned} E(X) &= E(u + \sigma z) \\ &= u + \sigma \cdot E(z) \\ &= u \end{aligned}$$

$$\begin{aligned} V(X) &= V(u + \sigma z) \\ &= \sigma^2 \cdot V(z) = \sigma^2 \end{aligned}$$

$$M_X(t) = E(e^{tx})$$

$$= E(e^{t(u + \sigma z)})$$

$$= e^{tu} \cdot E(e^{\sigma t z})$$

$$= e^{tu} \cdot e^{\frac{1}{2}(\sigma t)^2}$$

$$= e^{ut + \frac{1}{2}\sigma^2 t^2} \leftarrow$$

Additivity of $N(\mu, \sigma^2)$

Then:

$$X_i \sim N(\mu_i, \sigma_i^2).$$

X_1, \dots, X_n are indep.

$$X_1 + \dots + X_n \sim N\left(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2\right)$$

Def:

$$M_{X_i}(t) = e^{\frac{\mu_i t + \frac{1}{2} \sigma_i^2 t^2}{n}}$$

$$M_{X_1 + \dots + X_n}(t) = \prod_{i=1}^n M_{X_i}(t)$$

$$= e^{\sum_{i=1}^n (\mu_i t + \frac{1}{2} \sigma_i^2 t^2)}$$

$$= e^{\sum_{i=1}^n \mu_i t + \frac{1}{2} \sum_{i=1}^n \sigma_i^2 t^2}$$

This is the MGF of $N\left(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2\right)$

Example

$$X_1 \sim N(\mu_1, \sigma_1^2)$$

$$X_2 \sim N(\mu_2, \sigma_2^2)$$

$$X_1 + X_2 \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$X_1 - X_2 \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$$

Thm:

If $X_i \sim N(\mu_i, \sigma_i^2)$ indep.

then $\sum_{i=1}^n a_i X_i \sim N\left(\sum_{i=1}^n a_i \mu_i, \sum_{i=1}^n a_i^2 \sigma_i^2\right)$

Pf:

$$a_i X_i \sim N(a_i \mu_i, a_i^2 \sigma_i^2)$$

applying previous formula

Example:

$$X_1, \dots, X_n \stackrel{\text{IID}}{\sim} N(\mu, \sigma^2)$$

$$\bar{X} = \frac{X_1 + \dots + X_n}{n} = \sum_{i=1}^n a_i X_i$$

where $a_i = \frac{1}{n}$

$$\sum_{i=1}^n a_i \mu_i = \sum_{i=1}^n \left(\frac{1}{n} \cdot \mu \right) = \mu$$

$$\sum_{i=1}^n a_i^2 \sigma_i^2 = \sum_{i=1}^n \frac{1}{n^2} \cdot \sigma^2 = \frac{\sigma^2}{n}$$

so $\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$

